Regression Control Method with Stata

回归控制法及Stata应用

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Outline

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- 2. Model
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1. Introduction

- Regression control method (RCM)
 - Aka a panel data approach for program evaluation (Hsiao et al. 2012)
 - Exploits cross-sectional correlation to construct counterfactual outcomes for a single treated unit
 - Hsiao and Zhou (2019) propose to add covariates in model for helping prediction

1. Introduction

- Stata command rcm :
 - $\circ~$ Implement regression control method
 - $\circ~\mbox{Facilitate model selection}$ and estimation
 - Support to add covariates
 - Support placebo tests for statistical inferences

2. Model

2.1 Basic Model

- N cross-sectional units indexed by $i=1,\ldots,N$ over $t=1,\ldots,T_0+1,\ldots,T$ periods
- i=1 indexes the treated unit, whereas $\{2,\ldots,N\}$ is the index set of N-1 control units (donor pool)
- Policy intervention occurs at time $T_0+1,$ which partitions the time series into two sections:

Pre-treatment periods: $1, \ldots, T_0$ Post-treatment periods: $T_0 + 1, \ldots, T$

- y_{it}^1 and y_{it}^0 be the outcomes of the unit i in period t with and without intervention respectively
- y_{it} is observed that in the form:

$$y_{it} = d_{it} y_{it}^1 + \left(1 - d_{it}
ight) y_{it}^0$$

- $d_{it} = 1$ if i = 1 and unit i is under intervention in period t, or $d_{it} = 0$ if not
- The treatment effect for can be expressed as

$$\Delta_{it}=y^1_{it}-y^0_{it}$$
 $t=T_0+1,\ldots,T$

• Assume that y_{it}^0 is generated by a **pure** factor model of the form

$$y_{it}^0 = \mathbf{b}_i' \mathbf{f}_t + arepsilon_{it}$$

- $f_t {:} K imes 1$ (unobserved) common factors
- $\mathbf{b}_i': 1 imes K$ (unobserved) factor loadings
- ε_{it} : Random idiosyncratic component with $E(\varepsilon_{it}) = 0$.

Stack y_{it}^0 for $i \in \{2,\ldots,N\}$, we get

$$\widetilde{\mathbf{y}}_t = \widetilde{\mathbf{B}}\mathbf{f}_t + \widetilde{oldsymbol{arepsilon}}_t$$

•
$$\widetilde{\mathbf{y}}_t = \left(y_{2t}^0, \dots, y_{Nt}^0
ight)'$$

•
$$\widetilde{\boldsymbol{\varepsilon}}_t = (\varepsilon_{2t}, \dots, \varepsilon_{Nt})'$$

• $\widetilde{\mathbf{B}}: (N-1) imes K$ factor loading matrix $(\mathbf{b}_2, \dots, \mathbf{b}_N)'$

$$\mathbf{f}_t = \left(\widetilde{\mathbf{B}}' \widetilde{\mathbf{B}}
ight)^{-1} \widetilde{\mathbf{B}}' \left(\widetilde{\mathbf{y}}_t - \widetilde{oldsymbol{arepsilon}}_t
ight)$$

Then y_{1t}^0 can be represented as

$$egin{aligned} &y_{1t}^0 = \mathbf{b}_1' \mathbf{f}_t + arepsilon_{1t} \ &= \mathbf{b}_1' \left(\widetilde{\mathbf{B}}' \widetilde{\mathbf{B}}
ight)^{-1} \widetilde{\mathbf{B}}' \left(\widetilde{\mathbf{y}}_t - \widetilde{oldsymbol{arepsilon}}_t
ight) + arepsilon_{1t} \ &= oldsymbol{\gamma}' \widetilde{\mathbf{y}}_t + arepsilon_{1t} - oldsymbol{\gamma}' \widetilde{oldsymbol{arepsilon}}_t \end{aligned}$$

•
$$oldsymbol{\gamma}' = \mathbf{b}_1' \left(\widetilde{\mathbf{B}}' \widetilde{\mathbf{B}}
ight)^{-1} \widetilde{\mathbf{B}}'$$

• $\widetilde{\mathbf{y}}_t$ is correlated to $\varepsilon_{1t} - \boldsymbol{\gamma}' \widetilde{\boldsymbol{\varepsilon}}_t$

$$egin{aligned} y_{1t}^0 &= oldsymbol{\gamma}' \widetilde{oldsymbol{y}}_t + arepsilon_{1t} - oldsymbol{\gamma}' \widetilde{oldsymbol{arepsilon}}_t \ &= oldsymbol{\gamma}' \widetilde{oldsymbol{y}}_t + L\left(arepsilon_{1t} - oldsymbol{\gamma}' \widetilde{oldsymbol{arepsilon}}_t | \widetilde{oldsymbol{y}}_t
ight) + v_{1t} \ &= oldsymbol{\gamma}' \widetilde{oldsymbol{y}}_t + oldsymbol{c}_1 + oldsymbol{c}' \widetilde{oldsymbol{arepsilon}}_t + v_{1t} \end{aligned}$$

• $L(\varepsilon_{1t} - \gamma' \tilde{\varepsilon}_t | \tilde{\mathbf{y}}_t)$ is the linear projection of $\varepsilon_{1t} - \gamma' \tilde{\varepsilon}_t$ onto $(1, \tilde{\mathbf{y}}_t')$, and c_1 and \mathbf{c} are the minimizers of

$$\min_{c_1,\mathbf{c}} E\left[\left(arepsilon_{1t}-oldsymbol{\gamma}'\widetilde{oldsymbol{arepsilon}}_t-c_1-\mathbf{c}'\widetilde{oldsymbol{\mathrm{y}}}_t
ight)^2
ight]$$

$$egin{aligned} y_{1t}^0 &= oldsymbol{\gamma}' \widetilde{oldsymbol{y}}_t + c_1 + oldsymbol{c}' \widetilde{oldsymbol{y}}_t + v_{1t} \ &= \delta_1 + oldsymbol{\delta}' \widetilde{oldsymbol{y}}_t + v_{1t} \end{aligned}$$

•
$$\delta_1=c_1$$
 , $oldsymbol{\delta}'=oldsymbol{\gamma}'+\mathbf{c}'$

• Hsiao et al. (2012) advocate estimating $\hat{\delta}_1$ and $\hat{\delta}'$ by OLS with the pre-treatment subsample, and predicting the counterfactual outcomes as

$$\hat{y}_{1t}^{0} = \hat{\delta}_{1} + \hat{oldsymbol{\delta}}' \widetilde{\mathbf{y}}_{t}$$

- So the treatment effect Δ_{1t} can be predicted using

$$\hat{\Delta}_{1t} = y_{1t}^1 - \hat{y}_{1t}^0 \quad t = T_0 + 1, \dots, T$$

• Average treatment effect (ATE) is estimated by averaging $\hat{\Delta}_{1t}$ over the post-treatment periods:

$$\hat{\Delta}_1 = rac{1}{T-T_0}\sum_{t=T_0+1}^T \hat{\Delta}_{1t}$$

- Use all of control units for estimation may not be the best choice
- Large estimation variance in turn leads to poor out-of-sample predictions
- Hsiao et at. (2012) suggest using information criterion approach to select control units in exhaustive search (best subset selection)
- Use $\widetilde{\mathbf{y}}_t^*$ instead of $\widetilde{\mathbf{y}}_t$, where $\widetilde{\mathbf{y}}_t^*$ is the best subset of $(y_{2t},\ldots,y_{Nt})'$

2.2 Model with Covariates (Hsiao and Zhou, 2019)

• Assume y_{it}^0 is a function of p observable variables \mathbf{x}_{it} :

$$y_{it}^0 = \mathbf{x}_{it}^\prime eta + \mathbf{b}_i^\prime \mathbf{f}_t + arepsilon_{it}$$

• y_{1t}^0 can be predicted as

$$\hat{y}_{1t}^0 = \hat{\delta}_1 + \hat{oldsymbol{\delta}}' \mathbf{z}_t^*$$

• \mathbf{z}_t^* includes any subset of $\mathbf{z}_t = (y_{2t}, \dots, y_{Nt}, \mathbf{x}_{1t} \dots, \mathbf{x}_{Nt})'$ that helps to predict y_{1t}^0 (forward stepwise method or Lasso method suggested)

3 Extension

- Step 1 Select the Suboptimal Model Select the suboptimal model $M^*_{OLS}(j)$ using best subset selection, forward stepwise selection or backward stepwise selection, or $M^*_{lasso}(\lambda)$ using Lasso seletion $(j \in \{1, \ldots, T_0 - 1\})$
- Step 2 Select the Optimal Model Choose the optimal model M^{*} from all of suboptimal models in terms of information criterion or cross-validation
- After model selection, OLS or Lasso regression is used to fit the optimal model for counterfactual prediction

3.1 Select the Suboptimal Model

3.1.1 Best Subset Selection

- Hsiao et. al. (2012) suggest using best subset selection
- *exhaustion*:

For
$$k = 1, 2, \dots p$$
:
1. Fit all $\begin{pmatrix} p \\ k \end{pmatrix}$ models that contain k predictors using OLS regression

- 2. Choose the smallest RSS model as $M^{st}_{OLS}(k)$
- $2^p 1$ possible combination of the p predictors

3.1.1 Best Subset Selection

- *leaps and bounds* : quickly calculate best subsets without examining all possible subsets
- Fundamental inequality:

 $RSS(A) \leq RSS(B) \quad B \subset A$

- If $RSS(\{3,4\}) \geq RSS(1), RSS(\{3\})$ and $RSS(\{4\})$ do not need to be calculated

3.1.1 Best Subset Selection

- Initial : Reorder the variables by their impact on RSS
- Regression and bound tree (pair tree)
- Traverse all the subsets of the root node in level 1
- Traverse the tree from right to left

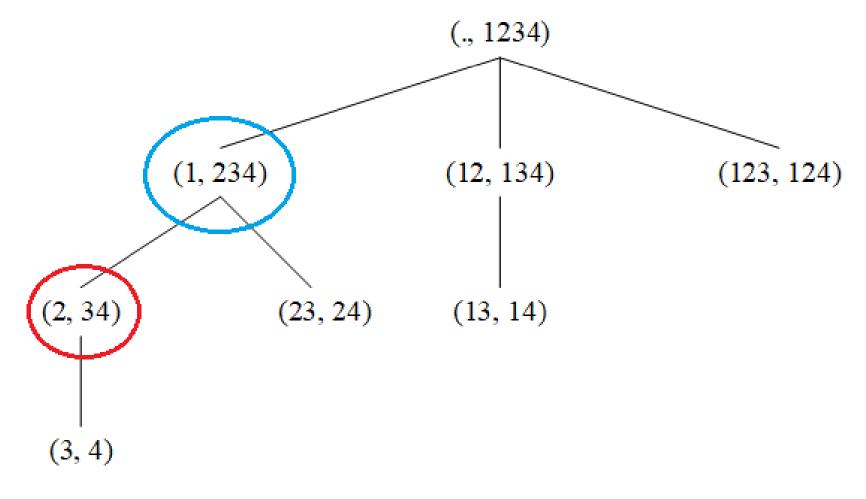


Figure : The pair tree

3.1.1 Best Subset Selection

- Suppose subset $\Omega \subset \{1,2,\ldots,p+1\}$, we have

$$\mathrm{RSS}(\Omega) = y'y - \left(\Phi'_\Omega \cdot y
ight)' \left(\Phi'_\Omega \Phi_\Omega
ight)^{-1} \left(\Phi'_\Omega \cdot y
ight)$$

- Φ is $T_0 imes (p+1)$ matrix with a constant column and predictors
- Precomputed matrix :

$$(y,\Phi)'(y,\Phi)=\left(egin{array}{cc} y'y & y'\Phi\ \Phi'y & \Phi'\Phi \end{array}
ight)$$

3.1.2 Forward Stepwise Selection

- Hsiao and Zhou (2019) suggest using forward stepwise selection
- $M^*_{OLS}(0)$ denote the smallest model, which contains no predictors
- For $k=0,\ldots,p-1$:
 - 1. Consider models that augment the predictors in $M^*_{OLS}(k)$ with one additional predictor, fit OLS regression
 - 2. Fit these models with OLS regression
 - 3. Choose the smallest RSS model as $M^{st}_{OLS}(k+1)$
- Shi and Huang (2021) : Forward-Selected PDA (fsPDa)

3.1.3 Backward Stepwise Selection

- $M^{\ast}_{OLS}(p)$ denote the largest possible model, which contains all p predictors
- For $k=p,\ldots,1$:
 - 1. Consider models that contain all but one of the predictors in $M^{\ast}_{OLS}(k)$
 - 2. Fit these models with OLS regression
 - 3. Choose the smallest RSS model as $M^{st}_{OLS}(k-1)$
- Require that $T_0 > p$ (the largest possible model can be fit)

3.1.4 Lasso Selection

• Lasso selects β to minimize:

$$\hat{oldsymbol{eta}}_{lasso}(\lambda) = rgmin_{oldsymbol{eta}} \sum_{i=1}^n \left(y_i - oldsymbol{x}_i^\prime oldsymbol{eta}
ight)^2 + \lambda \sum_{j=1}^p |eta_j|$$

- $\lambda \geq 0$ is a tuning parameter
- Two extreme cases : $\hat{m{eta}}_{lasso}(\lambda=0)=\hat{m{eta}}_{ols}, \hat{m{eta}}_{lasso}(\lambda=\infty)={m 0}$

3.1.4 Lasso Selection

- For $\lambda_l = \lambda_1, \dots, \lambda_L$:
 - 1. Calculate lasso estimator $\hat{oldsymbol{eta}}_{lasso}(\lambda_l)$
 - 2. Consider the model with a set of predictors $\Omega(\lambda_l)$, where

$$\Omega(\lambda_l) = \{predictor_j | j \in \{1, \dots, p\}, eta_{lasso, j}(\lambda_l)
eq 0\}$$

3. Choose the model as $M^*_{lasso}(\lambda_l)$ if $\Omega(\lambda_l)$ is different from $\Omega(\lambda_1),\ldots,\Omega(\lambda_{l-1})$

3.2 Select Optimal Model

3.2.1 Information criterion

$$egin{aligned} ext{AIC}(p) &= T_0 \ln \left(rac{\mathbf{e}_0' \mathbf{e}_0}{T_0}
ight) + 2(p+2) \ ext{BIC}(p) &= T_0 \ln \left(rac{\mathbf{e}_0' \mathbf{e}_0}{T_0}
ight) + (p+2) \ln(T_0) \ ext{AICc}(p) &= ext{AIC}(p) + rac{2(p+2)(p+3)}{T_1 - (p+1) - 2} \end{aligned}$$

• \mathbf{e}'_0 : OLS or Lasso residuals fitted in pre-treatment periods.

3.2.2 Cross-Validation

- Common practice after Lasso selection
- Optimize the out-of-sample prediction performance
- The mean squared error for each fold is computed as

$$ext{MSE}(\lambda,k) = rac{1}{n_k} \sum_{i \in \mathcal{F}_k} \left(y_i - oldsymbol{x}_i^\prime \hat{oldsymbol{eta}}_{lasso,k}(\lambda)
ight)^2$$

- K : Number of groups in which pre-treatment data is splitted
- n_k : Size of pre-treatment data partition k for k=1,...,K
- \mathcal{F}_k : Set of observations in k-fold

3.2.2 Cross-Validation

• K-fold Cross-Validation estimate of \mathbf{CV} **MSE**, which serves as a measure of prediction performance, is

$$egin{aligned} ext{CV MSE}(\lambda,K) =& rac{1}{K} \sum_{k=1}^K ext{MSE}(\lambda,k) \ &= & rac{1}{K} \sum_{k=1}^K \left(rac{1}{n_k} \sum_{i \in \mathcal{F}_k} \left(y_i - oldsymbol{x}_i^\prime \hat{oldsymbol{eta}}_{lasso,k}(\lambda)
ight)
ight) \end{aligned}$$

3.3 Post-Estimation of the Optimal Model

- Post-Estimation OLS :
 - Common practice
 - $\circ\,$ Fit OLS regression to the optimal model and obtain eta_{ols}
 - $\,\circ\,$ Use β_{ols} for counterfactual prediction
 - **Post-Lasso OLS** : Fit OLS regression after Lasso selected
- Post-Estimation Lasso:
 - $\circ\,$ Obtain $\beta_{lasso}(\lambda)$ in Lasso selection process
 - \circ Use $eta_{lasso}(\lambda)$ for counterfactual prediction
 - Can only be used after Lasso selection

3.4 Placebo Test

3.4.1 Placebo Test Using Fake Treatment Unit

- Reassign the treatment to control units (donor pool) where no intervention actually occurred
- Determine statistical significance of treatment effect
- p-val(t): p-value of estimated effect for a particular period is

$$p extstyle extstyle p extstyle val(t) = rac{1}{N}\sum_{i=1}^N \mathbb{1}\left(\left| \hat{\Delta}_{it}
ight| \geq \left| \hat{\Delta}_{1t}
ight|
ight) \quad t = T_0 + 1, \dots, T$$

3.4.1 Placebo Test Using Fake Treatment Unit

• *p-val*: The probability of obtaining a post/pre-MSPE ratio as large as that of treated unit, is

$$p\text{-}val = rac{1}{N}\sum_{i=1}^{N} 1\left(rac{ ext{MSPE}_{i,post}}{ ext{MSPE}_{i,pre}} \geq rac{ ext{MSPE}_{1,post}}{ ext{MSPE}_{1,pre}}
ight)$$

- Cutoff : Discard the fake units $i \in \{2, \dots, N\}$ with extreme values of $\mathrm{MSPE}_{i,pre}$ (Abadie et al. , 2010)

3.4.2 Placebo Test Using Fake Treatment Time

- Reassign the treatment to periods previous to the intervention when no treatment actually ocurred
- Whether there is a perceivable effect durning $1, \ldots, T_0 + 1$ periods?

4 The rcm command

- Implement "Regression Control Method (RCM)" in Stata
- Install: ssc install rcm, all replace with Stata version >= 16
- Installed files :
 - rcm.ado and rcm.sthlp: Stata ado and help file
- Ancillary files :
 - growth.dta : The dataset obtained from Hsiao et al. (2012) which has been reshaped to long form
 - repgermany.dta : The dataset obtained from Abadie et al. (2015) of which panel variable has been reencoded

4 The rcm command

• Syntax:

rcm depvar [indepvars] , trunit(#) trperiod(#) [options]

- xtset panelvar timevar must be used to declare a panel dataset in the usual long form
- depvar and indepvars must be numeric variables, and abbreviations are not allowed.
- Without indepvars : Basic model (Section 2.1)
- With indepvars : Model with covariates (Section 2.2)

5 Examples

5.1 Replicate Hsiao et al.(2012)

- Consider the impact on Hong Kong real GDP growth rate with the reversion of sovereignty on 1 July 1997 from the UK to China
 - Treatment period : 1997Q3
 - Pre-treatment preiods : 1993Q1-1997Q2
 - Post-treatment preiods : 1997Q3-2004Q1
 - Treated unit : Hong Kong
 - Control units : 10 countries/regions that are either in the region or economically closely associated with Hong Kong

5.1 Replicate Hsiao et al.(2012)

use growth, clear
xtset region time
des

| Observations: Variables: | | 1,525 3 | | 16 Jun 2019 00:03 |
|-----------------------------|------------------------|------------------------|----------------|-------------------|
| Variable name | Storage type | Display format | Value label | Variable label |
| time gdp region | float float long | %tq %8.0g %13.0g | region | |

Sorted by: region time

5.1 Replicate Hsiao et al.(2012)

label list

region:

- 1 Australia
- 2 Austria
- 3 Canada
- 4 China
- 5 Denmark
- 6 Finland
- 7 France
- 8 Germany
- 9 HongKong
- 10 Indonesia
- 11 Italy
- 12 Japan
- 13 Korea

- 14 Malaysia
- 15 Mexico
 - 16 Netherlands
- 17 NewZealand
- 18 Norway
- 19 Philippines
 - 20 Singapore
 - 21 Switzerland
 - 22 Taiwan
 - 23 Thailand
 - 24 UnitedKingdom
 - 25 UnitedStates

rcm gdp, trunit(9) trperiod(150) ctrlunit(4 10 12 13 14 19 20 22 23 25) postperiod(150/175)

- gdp : Specifies "gdp" as dependent variable (outcome variable)
- trunit(9) : Specifies "HongKong" as the treated unit
- trperiod(150) : Specifies "1997q3" as the treatment period
 (150 is obtained from di tq(1997q3)
- ctrlunit(4 10 12 13 14 19 20 22 23 25) : Specifies 10 countries/regions as the control units
- postperiod(150/175) : Specifies "1997q3-2003q4" as the posttreatment periods (175 is obtained from di tq(2003q4))

Selecting the suboptimal model with number of predictors 1-10...

Step 2: Select the optimal model from the suboptimal models
(criterion aicc specified)

Comparing the suboptimal models containing different set of predictors:

| K | AICc | AIC | BIC | R-squared |
|----|-----------|-----------|-----------|-----------|
| 1 | -144.7514 | -146.4657 | -143.7946 | 0.4034 |
| 2 | -160.5063 | -163.5832 | -160.0217 | 0.7937 |
| 3 | -170.6492 | -175.6492 | -171.1973 | 0.9056 |
| 4 | -171.7725 | -179.4088 | -174.0666 | 0.9314 |
| 5 | -169.7878 | -180.9878 | -174.7552 | 0.9438 |
| 6 | -164.2937 | -180.2937 | -173.1707 | 0.9477 |
| 7 | -156.6834 | -179.1834 | -171.1701 | 0.9503 |
| 8 | -146.2921 | -177.7207 | -168.8169 | 0.9517 |
| 9 | -131.7464 | -175.7464 | -165.9523 | 0.9518 |
| 10 | -111.3603 | -173.7603 | -163.0758 | 0.9518 |
| | | | | |

Among models with 1-10 predictors, the optimal model contains 4 predictors with AICc = -171.7725.

Fitting results in the pre-treatment periods using OLS:

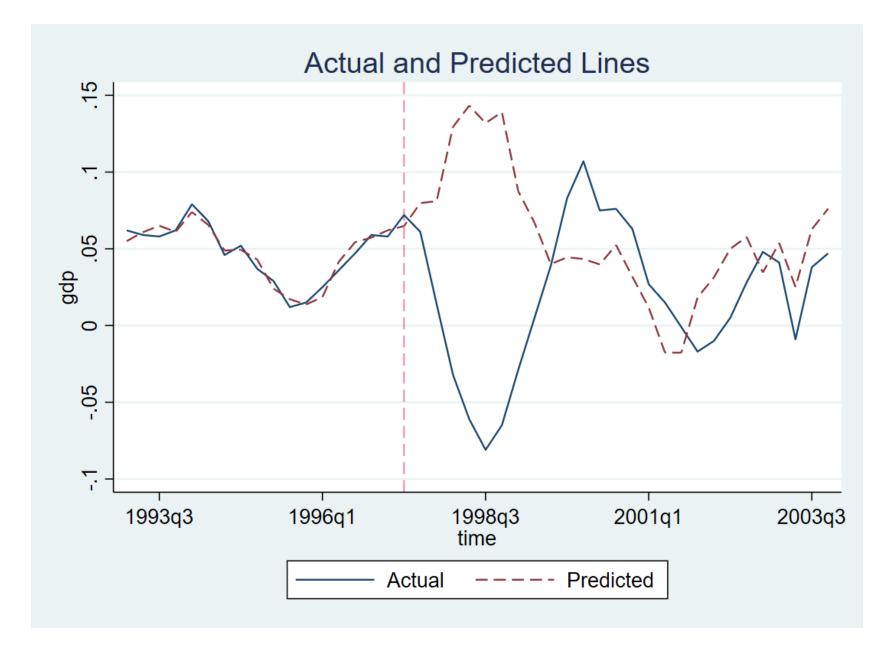
| Mean Absolute Err Mean Squared Erro Root Mean Squared | or = 0 | .00611 .00003 .00578 | Numbe | er of Obse er of Prec uared | | 18 4 0.93144 |
|---|--|--|--|---|---|--|
| gdp•HongKong | Coefficient | Std. err. | t | P> t | [95% conf. | interval] |
| gdp∙Korea gdp∙Japan gdp∙Taiwan gdp∙UnitedStates _cons | -0.4323 -0.6760 0.7926 0.4860 0.0263 | 0.0634 0.1117 0.3099 0.2195 0.0170 | -6.82 -6.05 2.56 2.21 1.54 | 0.000 0.000 0.024 0.045 0.147 | -0.5692 -0.9172 0.1231 0.0118 -0.0105 | -0.2954 -0.4347 1.4621 0.9603 0.0631 |

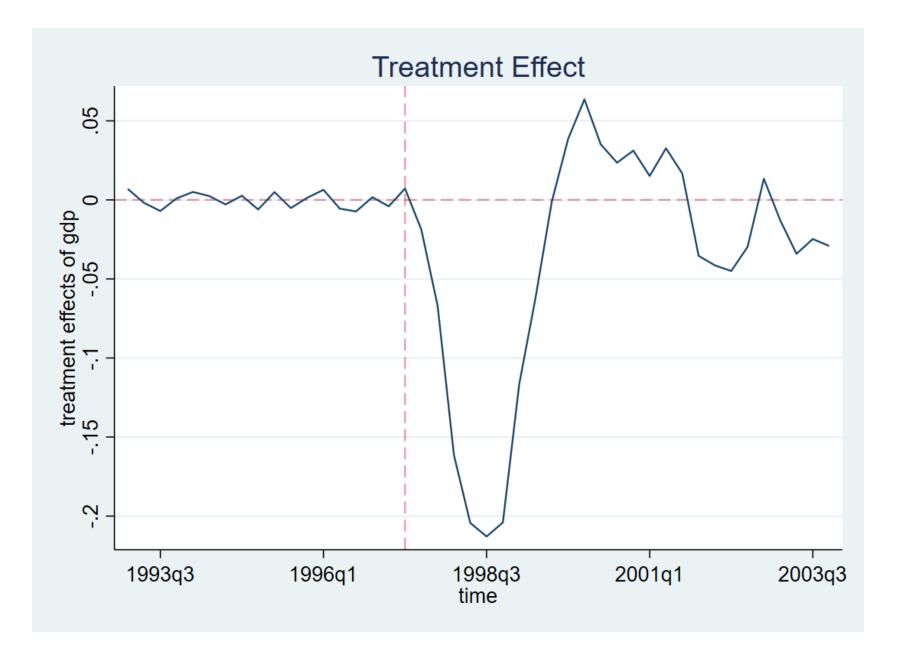
Prediction results in the post-treatment periods using OLS:

5.1 Replicate Hsiao et al.(2012)

| Time | Treated | Predicted | Tr. Effect |
|--------|---------|-----------|------------|
| 1997q3 | 0.0610 | 0.0798 | -0.0188 |
| 1997q4 | 0.0140 | 0.0810 | -0.0670 |
| 1998q1 | -0.0320 | 0.1294 | -0.1614 |
| 1998q2 | -0.0610 | 0.1433 | -0.2043 |
| 1998q3 | -0.0810 | 0.1319 | -0.2129 |
| 1998q4 | -0.0650 | 0.1390 | -0.2040 |
| 1999q1 | -0.0290 | 0.0876 | -0.1166 |
| 1999q2 | 0.0050 | 0.0670 | -0.0620 |
| 1999q3 | 0.0390 | 0.0400 | -0.0010 |
| 1999q4 | 0.0830 | 0.0445 | 0.0385 |
| 2000q1 | 0.1070 | 0.0434 | 0.0636 |
| 2000q2 | 0.0750 | 0.0398 | 0.0352 |
| 2000q3 | 0.0760 | 0.0524 | 0.0236 |
| 2000q4 | 0.0630 | 0.0318 | 0.0312 |
| 2001q1 | 0.0270 | 0.0118 | 0.0152 |
| 2001q2 | 0.0150 | -0.0177 | 0.0327 |
| 2001q3 | -0.0010 | -0.0177 | 0.0167 |
| 2001q4 | -0.0170 | 0.0184 | -0.0354 |
| 2002q1 | -0.0100 | 0.0314 | -0.0414 |
| 2002q2 | 0.0050 | 0.0500 | -0.0450 |
| 2002q3 | 0.0280 | 0.0577 | -0.0297 |
| 2002q4 | 0.0480 | 0.0346 | 0.0134 |
| 2003q1 | 0.0410 | 0.0538 | -0.0128 |
| 2003q2 | -0.0090 | 0.0251 | -0.0341 |
| 2003q3 | 0.0380 | 0.0628 | -0.0248 |
| 2003q4 | 0.0470 | 0.0761 | -0.0291 |
| Mean | 0.0180 | 0.0576 | -0.0396 |

Note: The average treatment effect over the post-treatment periods is -0.0396.





- Consider the impact on Hong Kong real GDP growth rate with the implementation of CEPA starting in 2004Q1 between mainland China and Hong Kong.
 - Treatment period : 2004Q1
 - Pre-treatment preiods : 1993Q1-2003Q4
 - Post-treatment preiods : 2004Q1-2008Q1
 - Treated unit : Hong Kong
 - Control units : All countries/regions in dataset except Hong Kong

rcm gdp, trunit(9) trperiod(176)

trperiod(176) : Specifies "2004q1" as the treatment period
 (176 is obtained from di tq(2004q1))

Step 1: Select the suboptimal models (method **best** specified) Note: If this takes too long, you may wish to try **method(lasso**)(recommended), **method(forward)** or **method(backward)**. Alternatively, you may restrict *indepvars*, and/or the donor pool by the option **ctrlunit()**.

Selecting the suboptimal model with number of predictors 1-24...

Step 2: Select the optimal model from the suboptimal models
(criterion aicc specified)

Comparing the suboptimal models containing different set of predictors:

| K | AICc | AIC | BIC | R-squared |
|----|-----------|-----------|-----------|-----------|
| 1 | -313.8269 | -314.4269 | -309.0743 | 0.5877 |
| 2 | -335.2386 | -336.2642 | -329.1275 | 0.7602 |
| 3 | -348.2800 | -349.8590 | -340.9380 | 0.8318 |
| 4 | -365.6420 | -367.9122 | -357.2071 | 0.8933 |
| 5 | -377.4412 | -380.5523 | -368.0630 | 0.9235 |
| 6 | -378.9426 | -383.0569 | -368.7833 | 0.9310 |
| 7 | -378.9074 | -384.2016 | -368.1439 | 0.9357 |
| 8 | -378.5854 | -385.2521 | -367.4102 | 0.9400 |
| 9 | -377.5003 | -385.7503 | -366.1242 | 0.9433 |
| 10 | -375.0098 | -385.0744 | -363.6641 | 0.9450 |
| 11 | -372.4606 | -384.5939 | -361.3994 | 0.9469 |
| 12 | -369.2578 | -383.7405 | -358.7619 | 0.9483 |
| 13 | -365.9158 | -383.0586 | -356.2958 | 0.9498 |
| 14 | -362.5660 | -382.7142 | -354.1671 | 0.9516 |
| 15 | -358.3157 | -381.8542 | -351.5230 | 0.9529 |
| 16 | -353.3736 | -380.7336 | -348.6182 | 0.9538 |
| 17 | -348.1579 | -379.8246 | -345.9250 | 0.9549 |
| 18 | -342.4931 | -379.0149 | -343.3311 | 0.9561 |
| 19 | -335.8492 | -377.8492 | -340.3812 | 0.9570 |
| 20 | -328.0881 | -376.2785 | -337.0264 | 0.9574 |
| 21 | -319.2286 | -374.4286 | -333.3922 | 0.9575 |
| 22 | -309.3373 | -372.4952 | -329.6747 | 0.9576 |
| 23 | -298.3113 | -370.5335 | -325.9288 | 0.9576 |
| 24 | -285.9617 | -368.5499 | -322.1610 | 0.9576 |
| | | | | |

Among models with 1-24 predictors, the optimal model contains 6 predictors with AICc = -378.9426.

Fitting results in the pre-treatment periods using OLS:

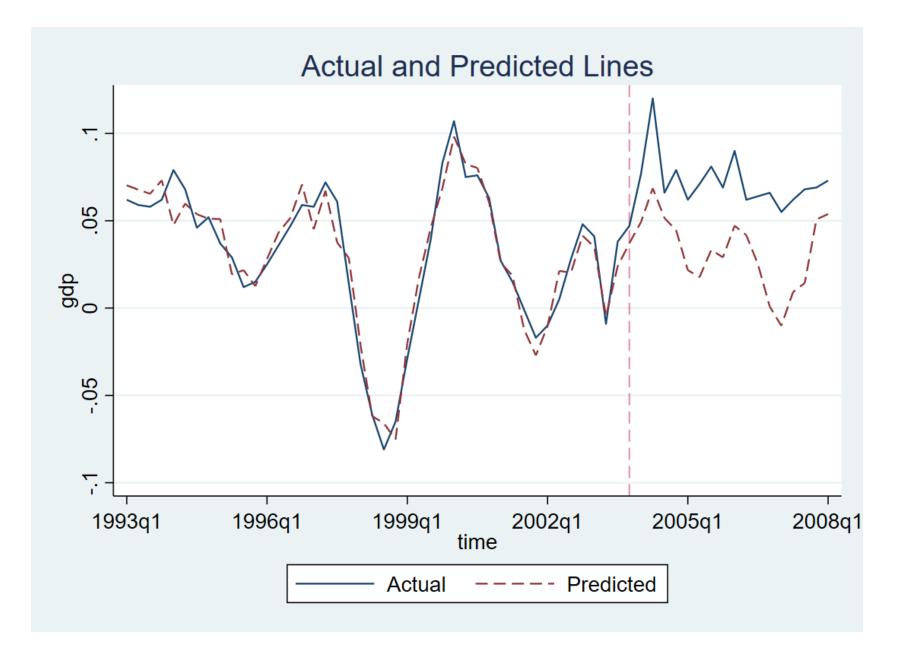
| Mean Absolute Error | = | 0.01070 | Number of Observations | = | 44 |
|-------------------------|---|---------|------------------------|---|---------|
| Mean Squared Error | = | 0.00014 | Number of Predictors | = | 6 |
| Root Mean Squared Error | = | 0.01170 | R-squared | = | 0.93097 |

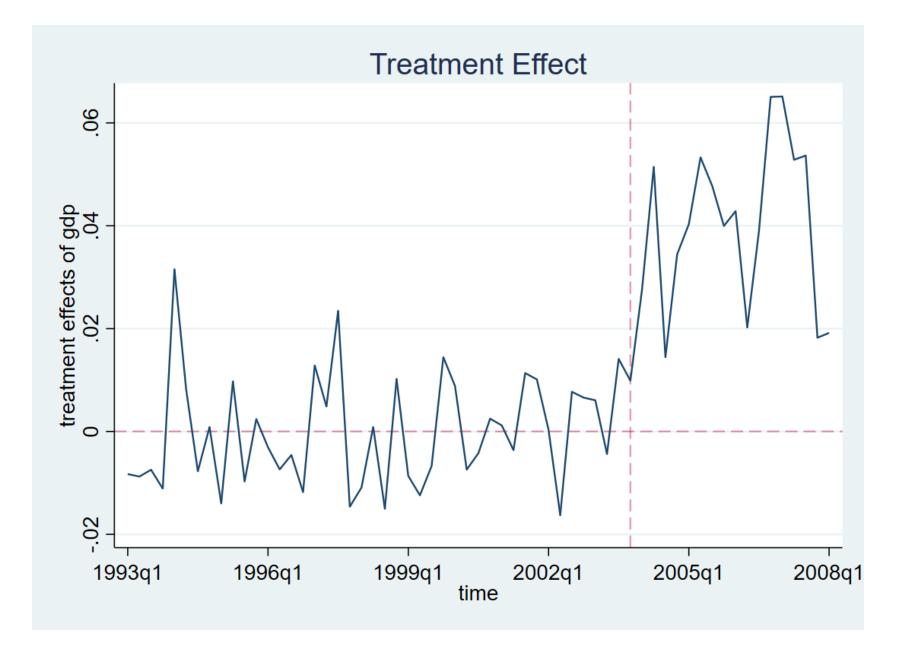
| gdp•HongKong | Coefficient | Std. err. | t | P> t | [95% conf. | interval] |
|---------------|-------------|-----------|-------|-------|------------|-----------|
| gdp•Norway | 0.3222 | 0.0538 | 5.99 | 0.000 | 0.2132 | 0.4311 |
| gdp.Austria | -1.0115 | 0.1682 | -6.01 | 0.000 | -1.3524 | -0.6707 |
| gdp•Korea | 0.3447 | 0.0469 | 7.35 | 0.000 | 0.2497 | 0.4398 |
| gdp.Mexico | 0.3129 | 0.0510 | 6.13 | 0.000 | 0.2095 | 0.4162 |
| gdp.Italy | -0.3177 | 0.1591 | -2.00 | 0.053 | -0.6400 | 0.0046 |
| gdp.Singapore | 0.1845 | 0.0546 | 3.38 | 0.002 | 0.0739 | 0.2951 |
| _cons | -0.0019 | 0.0037 | -0.52 | 0.603 | -0.0094 | 0.0056 |

Prediction results in the post-treatment periods using OLS:

| Time | Treated | Predicted | Tr. Effect |
|--------|---------|-----------|------------|
| 2004q1 | 0.0770 | 0.0493 | 0.0277 |
| 2004q2 | 0.1200 | 0.0686 | 0.0514 |
| 2004q3 | 0.0660 | 0.0515 | 0.0145 |
| 2004q4 | 0.0790 | 0.0446 | 0.0344 |
| 2005q1 | 0.0620 | 0.0217 | 0.0403 |
| 2005q2 | 0.0710 | 0.0177 | 0.0533 |
| 2005q3 | 0.0810 | 0.0333 | 0.0477 |
| 2005q4 | 0.0690 | 0.0290 | 0.0400 |
| 2006q1 | 0.0900 | 0.0471 | 0.0429 |
| 2006q2 | 0.0620 | 0.0417 | 0.0203 |
| 2006q3 | 0.0640 | 0.0250 | 0.0390 |
| 2006q4 | 0.0660 | 0.0009 | 0.0651 |
| 2007q1 | 0.0550 | -0.0101 | 0.0651 |
| 2007q2 | 0.0620 | 0.0092 | 0.0528 |
| 2007q3 | 0.0680 | 0.0143 | 0.0537 |
| 2007q4 | 0.0690 | 0.0508 | 0.0182 |
| 2008q1 | 0.0730 | 0.0538 | 0.0192 |
| Mean | 0.0726 | 0.0323 | 0.0403 |

Note: The average treatment effect over the post-treatment periods is 0.0403.





- Use the same dataset as Abadie et al. (2015)
- Estimate the economic impact of the 1990 German reunification
 - Treatment period : 1990
 - Pre-treatment preiods : 1960-1989
 - Post-treatment preiods : 1990-2003
 - Treated unit : West Germany
 - Control units : 16 OECD member countries

use repgermany.dta, clear
xtset country year
des

| Contains data from rep Observations: Variables: | | germany.dta 748 10 | I | 12 Aug 2021 08:25 |
|--|---------|--------------------------|---------|---------------------------------|
| Variable | Storage | Display | Value | Variable label |
| name | type | format | label | |
| year | float | %8.0g | | Year |
| gdp | long | %8.0g | | GDP per-capita (annual) |
| infrate | float | %9.0g | | Inflation Rate (annual) |
| trade | float | %9.0g | | Trade openness (annual) |
| schooling | float | %9.0g | | Schooling (every 5 years) |
| invest60 | float | %9.0g | | Investment rate (average 60-65) |
| invest70 | float | %9.0g | | Investment rate (average 70-75) |
| invest80 | float | %9.0g | | Investment rate (average 80-85) |
| industry | float | %9.0g | | Industry Share (annual) |
| country | long | %12.0g | country | Country Name |

Sorted by: country year

label list

country:

- 1 Australia
- 2 Austria
- 3 Belgium
- 4 Denmark
- 5 France
- 6 Greece
- 7 Italy
- 8 Japan
- 9 Netherlands
- 10 New Zealand
- 11 Norway
- 12 Portugal
- 13 Spain
- 14 Switzerland
- 15 UK
- 16 USA
- 17 West Germany

rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10)

- infrate trade industry : Specifies three covariates
- tru(17) : Abbr. for trunit(17)
- trp(1990) : Abbr.for trperiod(1990)
- me(lasso) : Abbr. for method(lasso), which specifies Lasso as selection method
- cr(cv) : Abbr. for criterion(cv), which specifies cross-validation as the selection criterion
- fold(10) : Specifies that cross-validation with 10 folds

Step 1: Select the suboptimal models
(method lasso specified)

Selecting the suboptimal model...

Step 2: Select the optimal model from the suboptimal models (criterion cv specified for **10-fold** cross-validation)

Comparing the suboptimal models containing different set of predictors:

| K | AICc | AIC | BIC | CV MSE | R-squared | lambda | Operation |
|---|----------|----------|----------|------------|-----------|-----------|---------------------------|
| 1 | 526.5586 | 525.6355 | 529.8391 | 2.187e+07 | 0.0888 | 4287.1880 | add gdp·Italy |
| 2 | 506.8945 | 505.2945 | 510.8993 | 1.038e+07 | 0.5668 | 2954.9924 | add gdp·Netherlands |
| 3 | 504.0514 | 501.5514 | 508.5574 | 8.616e+06 | 0.6403 | 2692.4790 | add gdp·Austria |
| 4 | 484.0258 | 480.3737 | 488.7808 | 4.061e+06 | 0.8289 | 1855.8213 | add gdp·Denmark |
| 5 | 478.7213 | 473.6304 | 483.4388 | 3.058e+06 | 0.8706 | 1614.0987 | add gdp·USA |
| 7 | 407.2748 | 398.2748 | 410.8855 | 2.420e+05 | 0.9894 | 459.7011 | add gdp·Greece gdp·Norway |
| 6 | 357.2235 | 350.3663 | 361.5759 | 58902.4925 | 0.9976 | 218.3953 | drop gdp·Netherlands |
| 7 | 334.6504 | 325.6504 | 338.2611 | 29433.5466 | 0.9988 | 150.5314 | add gdp·Switzerland |
| 8 | 328.6835 | 317.1046 | 331.1166 | 25005.4310 | 0.9990 | 137.1586 | add industry·Spain |
| 9 | 285.8035 | 271.1368 | 286.5500 | 6852.9127 | 0.9998 | 59.3727 | add gdp·Netherlands |
| 9 | 270.7571 | 256.0904 | 271.5036 | 4898.0417 | 0.9999 | 44.9133 | • |

Among models with 1-67 predictors, the optimal model contains 9 predictors with CV MSE = 4898.0417.

Fitting results in the pre-treatment periods using post-lasso OLS:

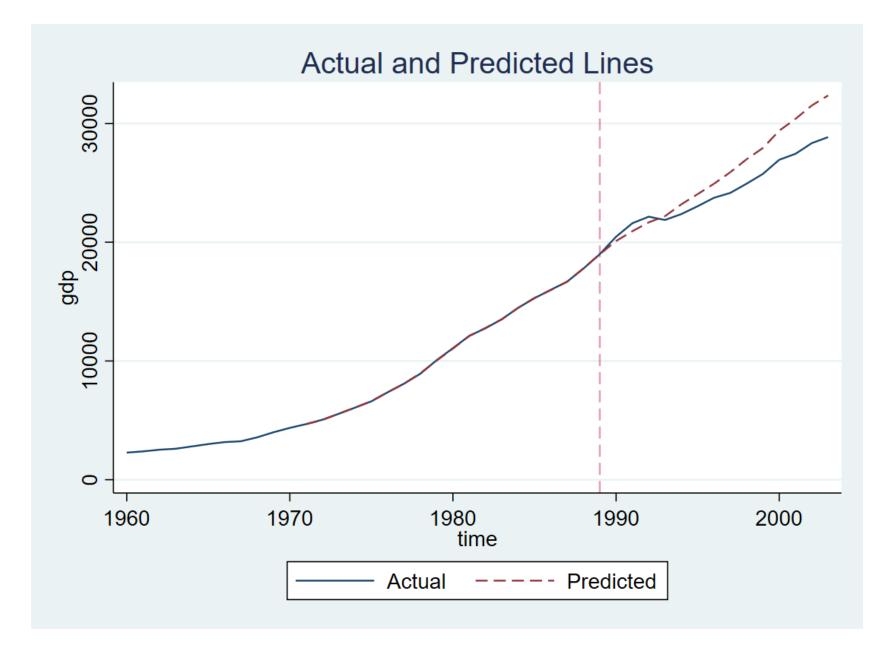
| Mean Absolute Error | = | 12.87119 | Number of Observations | = | 30 |
|-------------------------|---|----------|------------------------|---|---------|
| Mean Squared Error | = | 3.0e+02 | Number of Predictors | = | 9 |
| Root Mean Squared Error | = | 17.30368 | R-squared | = | 0.99999 |
| | | | | | |

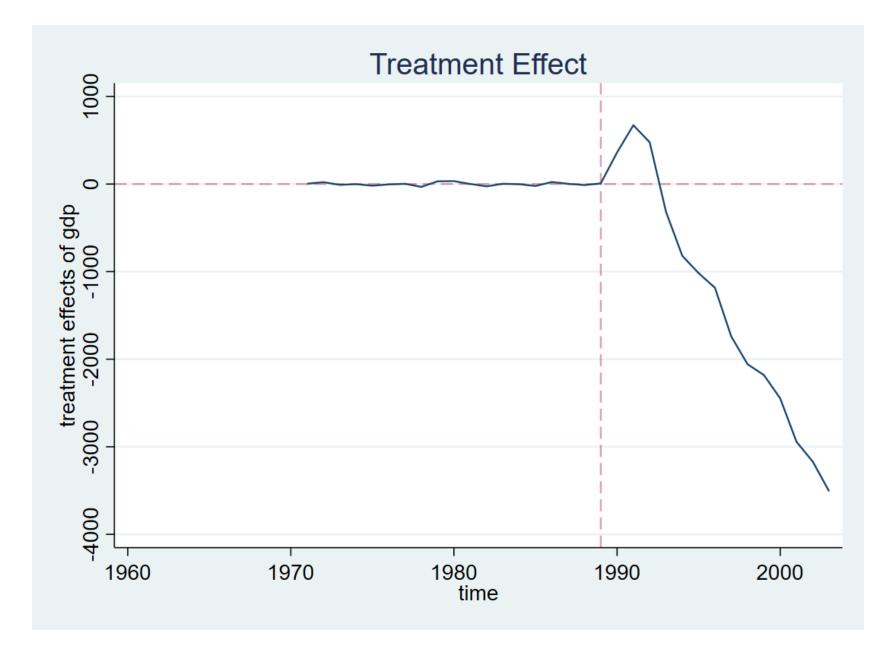
| gdp∙WestGermany | Coefficient | Std. err. | t | P> t | [95% conf. | interval] |
|-----------------|-------------|-----------|-------|-------|------------|-----------|
| gdp∙Austria | 0.1331 | 0.0708 | 1.88 | 0.093 | -0.0271 | 0.2934 |
| gdp.Denmark | 0.1046 | 0.0788 | 1.33 | 0.217 | -0.0735 | 0.2828 |
| gdp•Greece | 0.1513 | 0.0422 | 3.58 | 0.006 | 0.0557 | 0.2468 |
| gdp∙Italy | 0.2914 | 0.0838 | 3.48 | 0.007 | 0.1018 | 0.4809 |
| gdp•Netherlands | 0.1334 | 0.1097 | 1.22 | 0.255 | -0.1148 | 0.3816 |
| gdp.Norway | -0.0313 | 0.0598 | -0.52 | 0.613 | -0.1665 | 0.1039 |
| industry.Spain | -44.6926 | 11.6744 | -3.83 | 0.004 | -71.1019 | -18.2833 |
| gdp.Switzerland | 0.0548 | 0.0343 | 1.60 | 0.144 | -0.0228 | 0.1324 |
| gdp•USA | 0.2282 | 0.0439 | 5.20 | 0.001 | 0.1289 | 0.3276 |
| _cons | 1755.2090 | 419.7465 | 4.18 | 0.002 | 805.6764 | 2704.7417 |

Prediction results in the post-treatment periods using post-lasso OLS:

| Time | Treated | Predicted | Tr. Effect |
|------|------------|------------|------------|
| 1990 | 20465.0000 | 20104.6133 | 360.3867 |
| 1991 | 21602.0000 | 20930.3809 | 671.6191 |
| 1992 | 22154.0000 | 21677.2500 | 476.7500 |
| 1993 | 21878.0000 | 22194.6797 | -316.6797 |
| 1994 | 22371.0000 | 23190.9297 | -819.9297 |
| 1995 | 23035.0000 | 24052.6563 | -1017.6563 |
| 1996 | 23742.0000 | 24926.1309 | -1184.1309 |
| 1997 | 24156.0000 | 25896.0313 | -1740.0313 |
| 1998 | 24931.0000 | 26988.5430 | -2057.5430 |
| 1999 | 25755.0000 | 27935.7734 | -2180.7734 |
| 2000 | 26943.0000 | 29389.8184 | -2446.8184 |
| 2001 | 27449.0000 | 30392.2207 | -2943.2207 |
| 2002 | 28348.0000 | 31518.2871 | -3170.2871 |
| 2003 | 28855.0000 | 32363.5508 | -3508.5508 |
| Mean | 24406.0000 | 25825.7761 | -1419.7761 |

Note: The average treatment effect over the post-treatment periods is -1419.7761.





rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10) fill(mean) placebo(unit cut(10))

- fill(mean) : Fill in missing values by sample means for each units
- placebo(unit cut(10)):
 - Implement placebo tests using the fake treatment units
 - Discard the fake treatment units of which pre-treatment MSPE
 10 times smaller than or equal to that of the treated unit

Implementing placebo effects using fake treatment unit Australia...Austria...Belgium...Denmark...France...Greece...Italy...Japan...Netherlands...Ne
> wZealand...Norway...Portugal...Spain...Switzerland...UK...USA...

Placebo test results using fake treatment units:

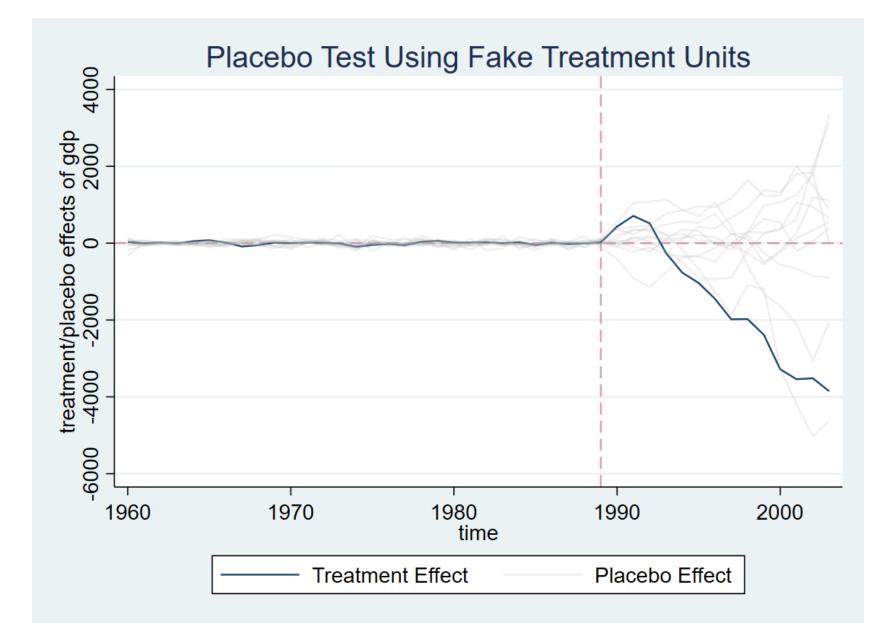
| Pre MSPE | Post MSPE | Post/Pre MSPE | Pre MSPE of Unit/Treated Uni |
|------------|---|---|---|
| 2009.4254 | 4912795.2300 | 2444.8757 | 1.0000 |
| 6616.8915 | 191860.4677 | 28,9956 | 3.2929 |
| 2640.2164 | 81819.2402 | 30.9896 | 1.3139 |
| 3509.5265 | 241028.1689 | 68.6783 | 1.7465 |
| 16088.6537 | 1218658.0663 | 75.7464 | 8.0066 |
| 2180.6256 | 181301.3038 | 83.1419 | 1.0852 |
| 10414.5510 | 2085867.0714 | 200.2839 | 5.1829 |
| 5532.6214 | 1278504.6038 | 231.0848 | 2.7533 |
| 7468.8217 | 858321.5550 | 114.9206 | 3.7169 |
| 8020.4273 | 1327015.9387 | 165.4545 | 3.9914 |
| 6950.5984 | 443053.1737 | 63.7432 | 3.4590 |
| 14405.5364 | 5967054.2037 | 414.2195 | 7.1690 |
| | 2009.4254 6616.8915 2640.2164 3509.5265 16088.6537 2180.6256 10414.5510 5532.6214 7468.8217 8020.4273 6950.5984 | 2009.4254 4912795.2300 6616.8915 191860.4677 2640.2164 81819.2402 3509.5265 241028.1689 16088.6537 1218658.0663 2180.6256 181301.3038 10414.5510 2085867.0714 5532.6214 1278504.6038 7468.8217 858321.5550 8020.4273 1327015.9387 6950.5984 443053.1737 | 2009.42544912795.23002444.87576616.8915191860.467728.99562640.216481819.240230.98963509.5265241028.168968.678316088.65371218658.066375.74642180.6256181301.303883.141910414.55102085867.0714200.28395532.62141278504.6038231.08487468.8217858321.5550114.92068020.42731327015.9387165.45456950.5984443053.173763.7432 |

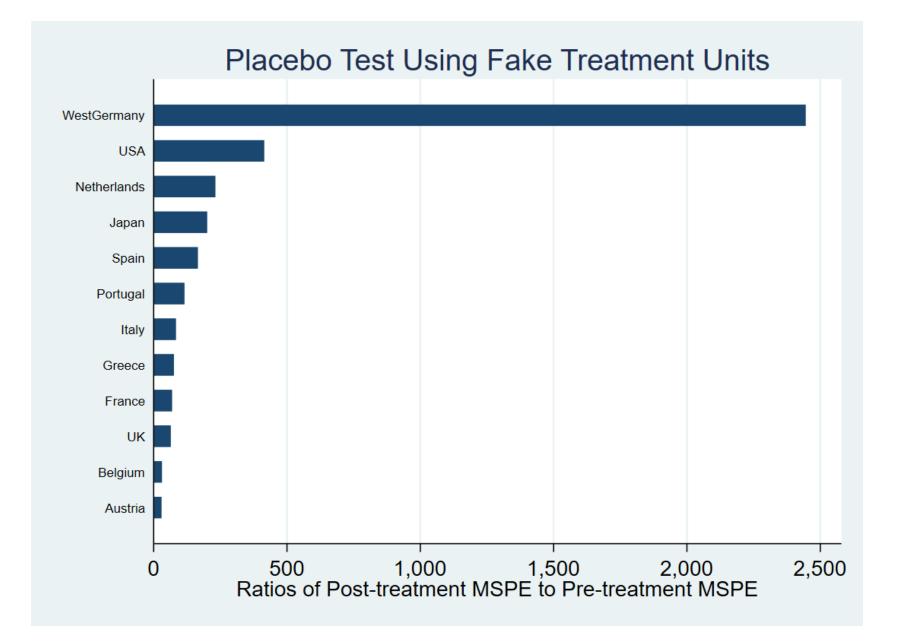
Note: The units Australia Denmark NewZealand Norway Switzerland with Pre-Treatment MSPE 10 times higher than WestGermany's are excluded. The probability of obtaining a post/pre-treatment MSPE ratio as large as WestGermany's is 0.0833.

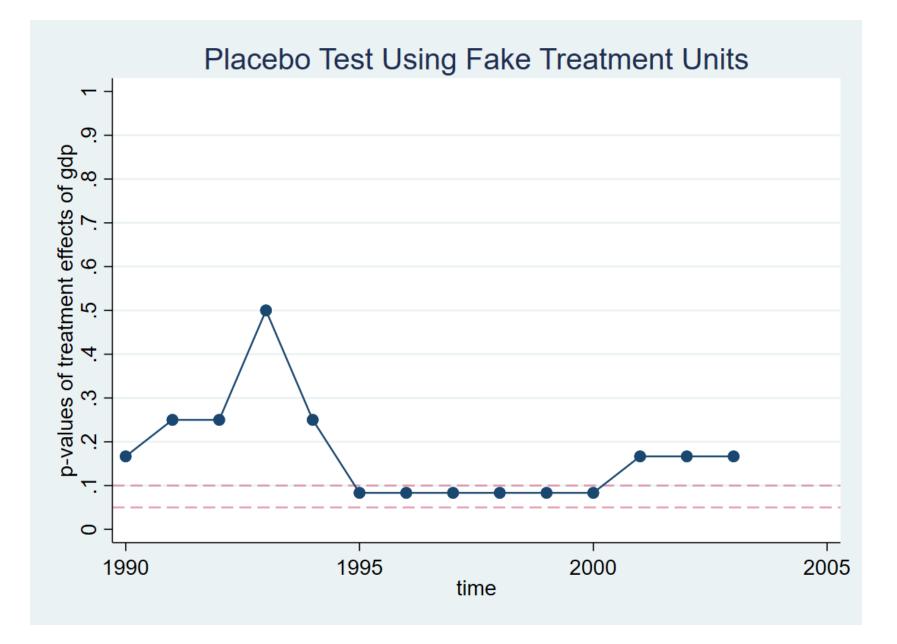
| Placebo test results | using fake | treatment units | (continued): |
|----------------------|------------|-----------------|--------------|
|----------------------|------------|-----------------|--------------|

| Time | Tr. Eff. | P-value |
|------|------------|---------|
| 1990 | 432.1543 | 0.1667 |
| 1991 | 709.9277 | 0.2500 |
| 1992 | 515.8379 | 0.2500 |
| 1993 | -256.7402 | 0.5000 |
| 1994 | -766.9336 | 0.2500 |
| 1995 | -1035.2422 | 0.0833 |
| 1996 | -1448.2324 | 0.0833 |
| 1997 | -1981.4160 | 0.0833 |
| 1998 | -1975.5234 | 0.0833 |
| 1999 | -2385.0527 | 0.0833 |
| 2000 | -3280.4258 | 0.0833 |
| 2001 | -3540.6426 | 0.1667 |
| 2002 | -3515.3242 | 0.1667 |
| 2003 | -3850.5918 | 0.1667 |
| | | |

Note: The p-value of the treatment effect for a particular period is definded as the frequency that the absolute values of the placebo effects are greater or equal to the absolute value of treatment effect.







rcm gdp infrate trade industry, tru(17) trp(1990) me(lasso) cr(cv) fold(10) fill(linear) placebo(period(1980 1985))

- fill(linear) : Fill in missing values by linear interpolation for each units
- placebo(period(1980 1985)) : Implement placebo tests with fake treatment time 1980 and 1985

Implementing placebo effects using fake treatment time 1980...1985...

5.2 Illustrate RCM with Covariates and Placebo Test

Placebo test results using fake treatment time 1980:

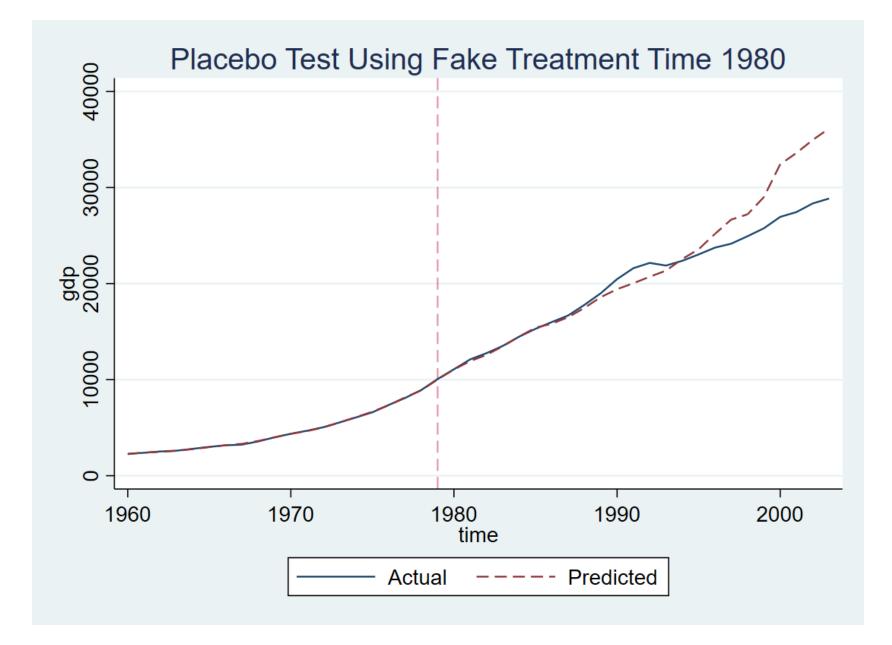
| Time | Treated | Predicted | Tr. Eff. |
|------|------------|------------|------------|
| 1980 | 11083.0000 | 11062.1953 | 20.8047 |
| 1981 | 12115.0000 | 11937.9014 | 177.0986 |
| 1982 | 12761.0000 | 12584.5449 | 176.4551 |
| 1983 | 13519.0000 | 13487.0322 | 31.9678 |
| 1984 | 14481.0000 | 14485.5996 | -4.5996 |
| 1985 | 15291.0000 | 15414.0498 | -123.0498 |
| 1986 | 15998.0000 | 15827.1543 | 170.8457 |
| 1987 | 16679.0000 | 16503.3066 | 175.6934 |
| 1988 | 17786.0000 | 17469.7988 | 316.2012 |
| 1989 | 18994.0000 | 18590.4004 | 403.5996 |
| 1990 | 20465.0000 | 19414.7520 | 1050.2480 |
| 1991 | 21602.0000 | 20036.1699 | 1565.8301 |
| 1992 | 22154.0000 | 20710.7656 | 1443.2344 |
| 1993 | 21878.0000 | 21346.8164 | 531.1836 |
| 1994 | 22371.0000 | 22564.2402 | -193.2402 |
| 1995 | 23035.0000 | 23573.0605 | -538.0605 |
| 1996 | 23742.0000 | 25172.0996 | -1430.0996 |
| 1997 | 24156.0000 | 26662.7520 | -2506.7520 |
| 1998 | 24931.0000 | 27212.3789 | -2281.3789 |
| 1999 | 25755.0000 | 28998.8145 | -3243.8145 |
| 2000 | 26943.0000 | 32415.6914 | -5472.6914 |
| 2001 | 27449.0000 | 33603.8672 | -6154.8672 |
| 2002 | 28348.0000 | 34962.1016 | -6614.1016 |
| 2003 | 28855.0000 | 36147.2969 | -7292.2969 |
| Mean | 20432.9583 | 21674.2829 | -1241.3246 |

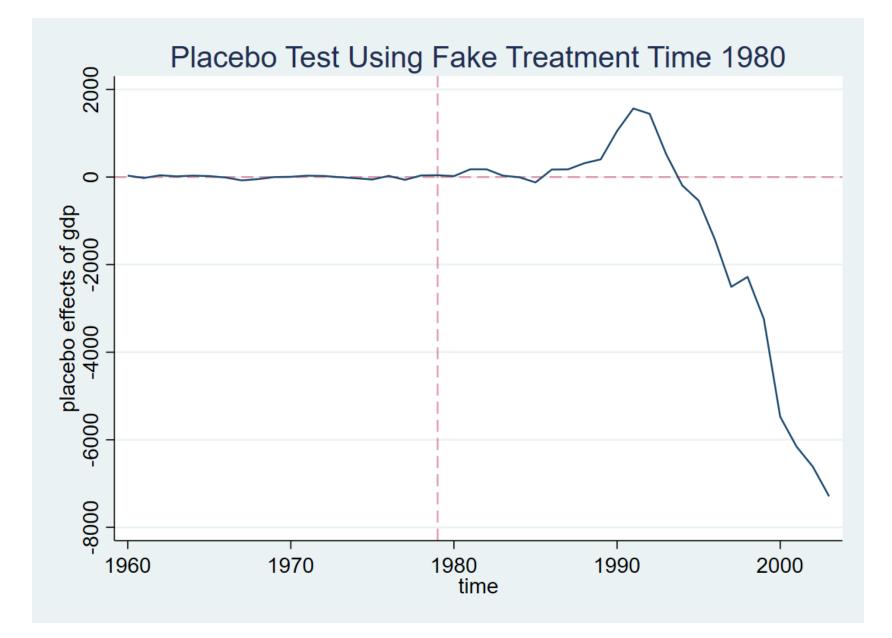
Note: The average treatment effect over the post-treatment periods is -1241.3246.

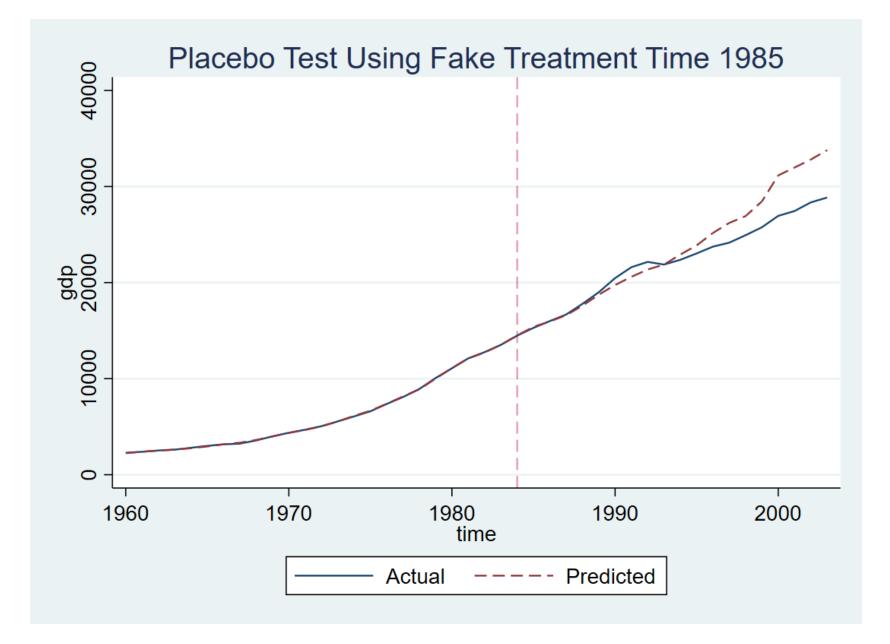
Placebo test results using fake treatment time 1985:

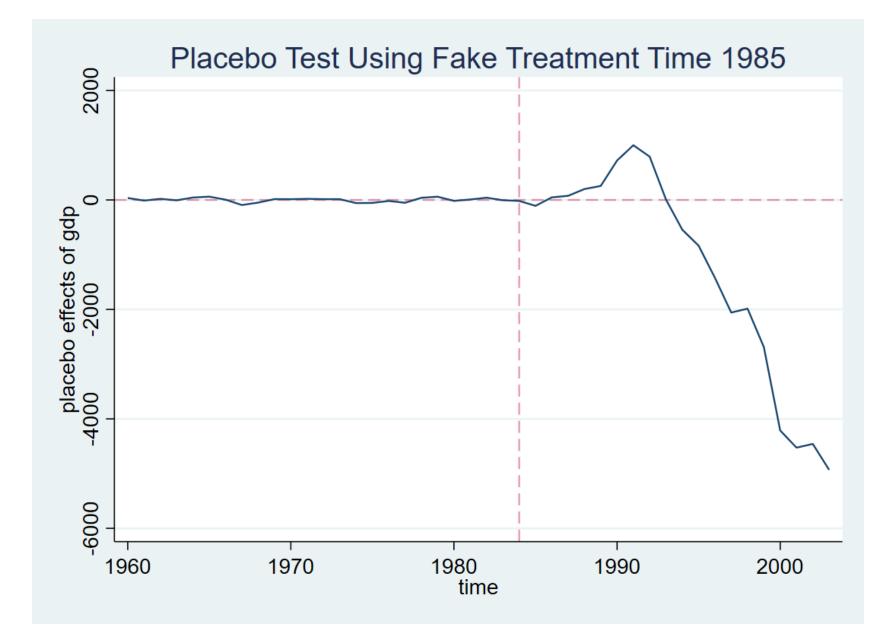
| Time | Treated | Predicted | Tr. Eff. |
|------|------------|------------|------------|
| 1985 | 15291.0000 | 15399.4453 | -108.4453 |
| 1986 | 15998.0000 | 15950.9844 | 47.0156 |
| 1987 | 16679.0000 | 16604.2344 | 74.7656 |
| 1988 | 17786.0000 | 17587.4297 | 198.5703 |
| 1989 | 18994.0000 | 18738.6387 | 255.3613 |
| 1990 | 20465.0000 | 19745.6875 | 719.3125 |
| 1991 | 21602.0000 | 20602.5215 | 999.4785 |
| 1992 | 22154.0000 | 21363.2012 | 790.7988 |
| 1993 | 21878.0000 | 21871.5781 | 6.4219 |
| 1994 | 22371.0000 | 22915.0332 | -544.0332 |
| 1995 | 23035.0000 | 23870.9434 | -835.9434 |
| 1996 | 23742.0000 | 25165.6582 | -1423.6582 |
| 1997 | 24156.0000 | 26213.7539 | -2057.7539 |
| 1998 | 24931.0000 | 26917.2168 | -1986.2168 |
| 1999 | 25755.0000 | 28444.0781 | -2689.0781 |
| 2000 | 26943.0000 | 31154.6328 | -4211.6328 |
| 2001 | 27449.0000 | 31974.5781 | -4525.5781 |
| 2002 | 28348.0000 | 32806.0039 | -4458.0039 |
| 2003 | 28855.0000 | 33786.4219 | -4931.4219 |
| Mean | 22443.7895 | 23742.7390 | -1298.9495 |

Note: The average treatment effect over the post-treatment periods is -1298.9495.









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Thank for Listening

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